Administrative Items: Homework 1

- Homework 1 Out Now (Due 9/7 @ 11:59pm)
- Counts as 6 points towards final grade
 - HWs are generally 6 or 7 points each for a total of 60%
- 4 Questions aimed at probability of random events
- Available on course website: <u>pachecoj.com/courses/csc380_fall21/</u>
- Submit PDF of answers + work on D2L

Administrative Items: Office Hours

Enfa

- Mondays, 10:30-11:30am (hybrid)
- In-person component: Gould-Simpson Rm 934, Desk #6

Saiful

- Tuesdays, 10-11am (hybrid)
- In-person component: Gould-Simpson Rm 942

Jason Wednesdays, 10-11am (Zoom only)

Zoom office hour meeting coordinates available on D2L

Recap

- > A random process is modeled by:
 - ightharpoonup Sample space Ω is the set of all possible outcomes
 - ightharpoonup **Events** E each being a subset of Ω
 - ightharpoonup Probability function P assigns a probability in [0,1] to each event
- > Axioms of probability
 - 1. For any event E, $0 \le P(E) \le 1$
 - 2. $P(\Omega) = 1$ and $P(\emptyset) = 0$
 - 3. For any *finite* or *countably infinite* sequence of pairwise mutually disjoint events E_1, E_2, E_3, \ldots

$$P\Big(\bigcup_{i\geq 1} E_i\Big) = \sum_{i\geq 1} P(E_i)$$

Recap

 \blacktriangleright A random variable is a <u>function</u> of samples to real values: $X:\Omega\to\mathbb{R}$

ightharpoonup X = x is an event with probability: $p(X = x) = \sum_{\omega \in \Omega: X(\omega) = x} P(\omega)$

- > Some fundamental rules of probability:
 - > Conditional: $p(X \mid Y) = \frac{p(X,Y)}{p(Y)} = \frac{p(X,Y)}{\sum_{x} p(X=x,Y)}$
 - \blacktriangleright Law of total probability: $p(Y) = \sum_{x} p(Y, X = x)$
 - ightharpoonup Probability chain rule: $p(X,Y)=p(Y)p(X\mid Y)$

Outline

- Useful Discrete Distributions (+numpy.random)
- Continuous Probability
- Useful Continuous Distributions

Outline

- ➤ Useful Discrete Distributions (+numpy.random)
- Continuous Probability
- > Useful Continuous Distributions

Numpy Library

Package containing many useful numerical functions... NumPy



CONDA

If you use conda, you can install NumPy from the defaults or conda-forge channels:

```
conda create -n my-env
conda activate my-env
conda config --env --add channels conda-forge
conda install numpy
```

PIP

If you use pip, you can install NumPy with:

```
pip install numpy
```

...we are interested in numpy.random at the moment



- Lightweight library for sampling random variables
- Supports most standard discrete PMFs and continuous PDFs
- Also handles random permutations of lists
- Imported along with Numpy as,

```
import numpy as np
```

- Functions accessible via np.random.functionname
- There are multiple random number generators... distinguishing them and seeding them can get a bit confusing...

Docs: https://numpy.org/doc/1.16/reference/routines.random.html

Allows sampling from many common distributions

Set (global) random seed as,

```
import numpy as np
seed = 12345
np.random.seed(seed)
```

- Fine for this class, but a bad habit to get into
- Sets global seed, which can be problematic in larger projects
- Better to create new instance of the Random Number Generator (RNG)

```
seed = 67890
rng = np.random.default_rng(seed)
```

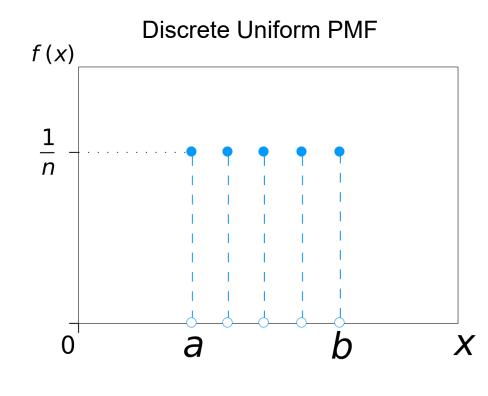
Pass around rng object to generate random numbers

Until now, we have worked with the **uniform distribution** on N values with PMF,

$$p(X = k) = \frac{1}{N}$$

More generally, we define on an interval [a,b] and assign a *named* PMF:

Uniform
$$(X = k; a, b, N) = \frac{1}{N}$$



The average or mean or expected value from the Uniform is:

$$E[X] = \frac{a+b}{2}$$

There are many other useful, but non-uniform, probability distributions

numpy.random.randint

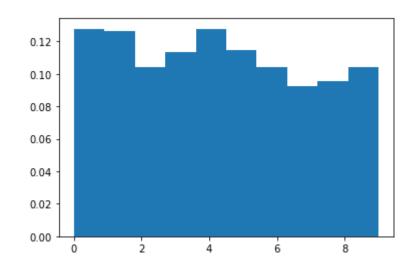
numpy.random.randint(low, high=None, size=None, dtype='l')

Return random integers from low (inclusive) to high (exclusive).

Return random integers from the "discrete uniform" distribution of the specified dtype in the "half-open" interval [low, high). If high is None (the default), then results are from [0, low).

Sample a discrete uniform random variable,

```
import matplotlib.pyplot as plt
X = np.random.randint(0,10,1000)
count, bins, ignored = plt.hist(X, 10, density=True)
plt.show()
```



- Caution Interval is [low,high) and upper bound is exclusive
- Most calls (but not all) in numpy involving intervals follow this pattern
- Size argument accepts tuples for sampling ndarrays

Bernoulli A.k.a. the **coinflip** distribution on <u>binary</u> RVs $X \in \{0,1\}$

$$p(X) = \pi^X (1 - \pi)^{(1 - X)}$$

Where π is the probability of **success** (e.g. heads), and also the mean

$$\mathbf{E}[X] = \pi \cdot 1 + (1 - \pi) \cdot 0 = \pi$$

Suppose we flip N independent coins X_1, X_2, \dots, X_N , what is the distribution over their sum $Y = \sum_{i=1}^{N} X_i$

Num. "successes" out of N trials
$$p(Y=k) = \binom{N}{k} \pi^k (1-\pi)^{N-k}$$
 Binomial Dist.

Binomial Mean: $\mathbf{E}[Y] = N \cdot \pi$ Sum of means for N indep. Bernoulli RVs



numpy.random.binomial

numpy.random.binomial(n, p, size=None)

Draw samples from a binomial distribution.

Samples are drawn from a binomial distribution with specified parameters, n trials and p probability of success where n an integer >= 0 and p is in the interval [0,1]. (n may be input as a float, but it is truncated to an integer in use)

Binomial PMF

 $p(Y=k)=inom{N}{k}\pi^k(1-\pi)^{N-k}$

Example A company drills 9 wild-cat oil exploration wells, each with an estimated probability of success of 0.1. All nine wells fail. What is the probability of that happening?

Answer this by simulating 20,000 trials...

```
N = 20000
p = 0.1
wells = 9
X = np.random.binomial(wells, p, N)
odds = sum( X == 0 )/N
odds
```



Question: How many flips until we observe a success?

Geometric Distribution on number of independent draws of $X \sim \text{Bernoulli}(\pi)$ until success:

$$p(Y=n)=(1-\pi)^{n-1}\pi$$

$$\mathbf{E}[Y]=rac{1}{\pi} \qquad \qquad \mathbf{E}[Y]=rac{1}{\pi} \qquad \qquad \mathbf{E}[Y]=rac{1}{\pi} \qquad \qquad \mathbf{E}[Y]=\frac{1}{\pi} \qquad \qquad \mathbf{E}[Y]=\frac{1}{$$

e.g. there must be n-1 failures (tails) before a success (heads).

Question: How many more flips of we have already seen k failures?

$$p(Y = n + k \mid Y > k) = \frac{p(Y = n + k, Y > k)}{p(Y > k)} = \frac{p(Y = n + k)}{p(Y > k)}$$
$$= \frac{(1 - \pi)^{n + k - 1} \pi}{\sum_{i = k}^{\infty} (1 - \pi)^{i} \pi} = \frac{(1 - \pi)^{n + k - 1} \pi}{(1 - \pi)^{k}} = (1 - \pi)^{n - 1} \pi = p(Y = n)$$

For
$$0 < x < 1, \sum_{i=k}^{\infty} x^i = x^k/(1-x)$$

Corollary: $p(Y > k) = (1 - \pi)^{k-1}$



Categorical Distribution on integer-valued RV $X \in \{1, \dots, K\}$ (

$$p(X) = \prod_{k=1}^K \pi_k^{\mathbf{I}(X=k)} \quad \text{or} \quad p(X) = \sum_{k=1}^K \mathbf{I}(X=k) \cdot \pi_k$$

with parameter $p(X = k) = \pi_k$ and Kronecker delta:

$$\mathbf{I}(X=k) = \begin{cases} 1, & \text{If } X = k \\ 0, & \text{Otherwise} \end{cases}$$

Can also represent X as one-hot binary vector,

$$X \in \{0,1\}^K$$
 where $\sum_{k=1}^K X_k = 1$ then $p(X) = \prod_{k=1}^K \pi_k^{X_k}$

This representation is special case of the multinomial distribution

What if we count outcomes of N independent categorical RVs?

Multinomial Distribution on K-vector $X \in \{0, N\}^K$ of counts of N repeated trials $\sum_{k=1}^K X_k = N$ with PMF:

$$p(x_1, \dots, x_K) = \binom{n}{x_1 x_2 \dots x_K} \prod_{k=1}^K \pi_k^{x_k}$$

Number of ways to partition N objects into K groups:

$$\binom{n}{x_1 x_2 \dots x_K} = \frac{n!}{x_1! x_2! \dots x_K!}$$

Leading term ensures PMF is properly normalized:

$$\sum_{x_1} \sum_{x_2} \dots \sum_{x_K} p(x_1, x_2, \dots, x_K) = 1$$

numpy.random.multinomial

numpy.random.multinomial(n, pvals, size=None)

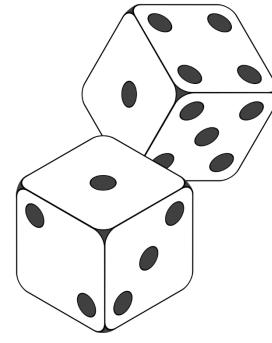
Draw samples from a multinomial distribution.

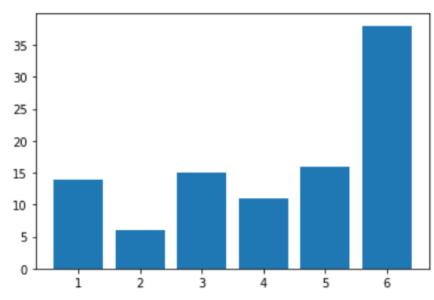
The multinomial distribution is a multivariate generalisation of the binomial distribution. Take an experiment with one of p possible outcomes. An example of such an experiment is throwing a dice, where the outcome can be 1 through 6. Each sample drawn from the distribution represents n such experiments. Its values, $X_i = [X_0, X_1, ..., X_p]$, represent the number of times the outcome was i.

Example Simulate 100 throws of a "loaded" die that has 3X the chance of rolling 6, and equal chance for remaining numbers.

```
N = 100
p_unnorm = np.array([1,1,1,1,1,3])
p = p_unnorm / sum(p_unnorm) # normalize
X = np.random.multinomial(N, p)
plt.bar(np.arange(6) + 1, X)
plt.show()
```

Note: Probability vector <u>must</u> be valid PMF (nonnegative, normalized a.k.a sum to 1)





How to simulate Bernoulli? Categorical?

Bernoulli is equivalent to a single draw from a binomial,

```
X = np.random.binomial(n=1, p=0.5) # fair coin flip
print(X)
```

Categorical is equivalent to a single draw from a multinomial,

```
X = np.random.multinomial(1, [0.5, 0.5]) # also a fair coin flip
print(X)
[0 1]
```

A **Poisson** RV X with <u>rate</u> parameter λ has

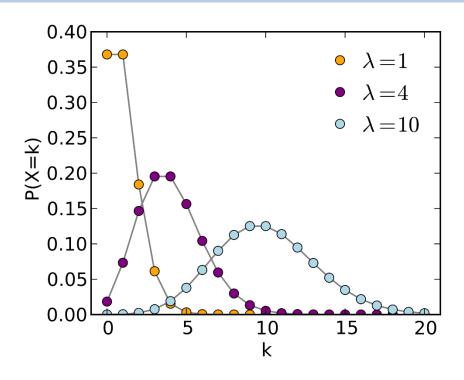
the following distribution:

$$p(X = k) = \frac{e^{-\lambda} \lambda^k}{k!}$$

Mean and variance both scale with parameter

$$\mathbf{E}[X] = \mathbf{Var}[X] = \lambda$$

Represents number of times an *event* occurs in an interval of time or space.



Ex. Probability of overflow floods in 100 years,

$$p(k \text{ overflow floods in } 100 \text{ yrs}) = \frac{e^{-1}1^k}{k!}$$

Avg. 1 overflow flood every 100 years, makes setting rate parameter easy.

Additivity The sum of a finite number of Poisson RVs is a Poisson RV.

$$X \sim \text{Poisson}(\lambda_1), \quad Y \sim \text{Poisson}(\lambda_2), \quad X + Y \sim \text{Poisson}(\lambda_1 + \lambda_2)$$

numpy.random.poisson

numpy.random.poisson(lam=1.0, size=None)

Draw samples from a Poisson distribution.

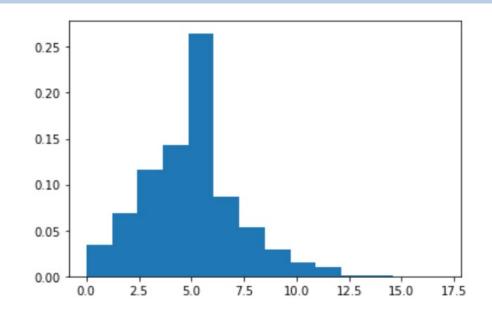
The Poisson distribution is the limit of the binomial distribution for large N.

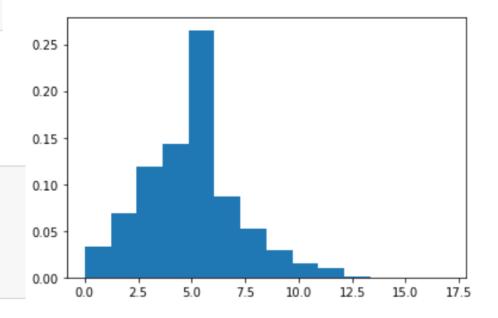
Example Simulate 100,000 draws from a Poisson with rate 5.0,

```
X = np.random.poisson(5, 100000)
ount, bins, ignored = plt.hist(X, 14, density=True)
plt.show()
```

Additivity The sum of a finite number of Poisson RVs is a Poisson RV.

```
X = np.random.poisson(2.5, 100000)
Y = np.random.poisson(2.5, 100000)
count, bins, ignored = plt.hist(X+Y, 14, density=True)
plt.show()
```





Administrative Items

- Special office hours
 - With Jason
 - Tomorrow @ 11am (Zoom)
 - See D2L Events for Zoom coordinates
 - TA office hours Monday (?) / Tuesday next week

Revision to HW1 (See yesterday's note on Piazza)

Problem 1(e) guidance...